# Gauss versus Laplace rates of contraction under Besov regularity

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### Joint work with

- S. Agapiou, M. Dashti and T. Helin, *Rates of contraction of posterior distributions based on p-exponential priors*, arXiv:1811.12244 (to appear in Bernoulli).
- S. Agapiou and A. Savva, *Adaptive rates of contraction based on p-exponential priors*, in preparation.
- S. Agapiou and S. Wang, *Frequentist rates of contraction in Bayesian inverse problems with Laplace priors*, in preparation.

## Outline



- WNM Minimax rates under Besov regularity
- p-exponential measures
- WNM ROC under Besov regularity

### 5 Numerics



## Outline



2 WNM - Minimax rates under Besov regularity

3 *p*-exponential measures

- WNM ROC under Besov regularity
- 5 Numerics



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Conclusion

# Multiscale features in images



WNM - ROC under Besov regularity

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## Wavelet expansions

•  $\{\psi_\ell\}_{\ell=1}^\infty$  orthonormal basis for  $L_2(\mathbb{T}^m)$ 

$$u(x) = \sum_{\ell=1}^{\infty} u_{\ell} \psi_{\ell}(x), \quad u_{\ell} = \langle u, \psi_{\ell} \rangle.$$

- $\bullet$  e.g.  $\{\psi_\ell\}$  is the Fourier basis
- For functions with multiscale features, better use wavelet bases  $\{\psi_{kl}\}$

$$u(x) = \sum_{k=1}^{\infty} \sum_{l=1}^{2^k} u_{kl} \psi_{kl}(x), \quad u_{kl} = \langle u, \psi_{kl} \rangle.$$
e.g. 2D Haar

### Besov Spaces

- Functions identified with expansion coefficients  $(u_\ell) \in \ell_2$  or  $(u_{kl}) \in \ell_2$
- Besov space of smoothness  $s \in \mathbb{R}$ , with integrability parameter  $q \ge 1$

$$B_{qq}^{s} = \left\{ u \in \mathbb{R}^{\infty} : \sum_{\ell=1}^{\infty} \ell^{q(\frac{s}{d} + \frac{1}{2}) - 1} |u_{\ell}|^{q} < \infty \right\}, \quad \|u\|_{B_{qq}^{s}} = \left( \sum_{\ell=1}^{\infty} \ell^{q(\frac{s}{d} + \frac{1}{2}) - 1} |u_{\ell}|^{q} \right)^{\frac{1}{q}}$$

• q = 2:  $B_{22}^s = H^s$ , Sobolev Hilbert spaces

- $q = \infty$ ,  $s \notin \mathbb{N}$ :  $B^s_{\infty\infty} = C^s$ , Hölder spaces
- Smaller q associated with sparsity and spatial inhomogeneity

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*p*-exponential measures

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## Besov Spaces - Intuition



I.M. Johnstone, *Gaussian estimation: sequence and wavelet models*, draft book.

### Function-space priors via random series expansions

$$u(x) = \sum_{\ell=1}^\infty u_\ell \psi_\ell(x)$$

• Randomize coefficients:  $u_{\ell} = \gamma_{\ell} \xi_{\ell}$  where  $\xi_{\ell} \stackrel{iid}{\sim} f$ ,  $\gamma_{\ell} > 0$  decaying scalings

- Choice of wavelet basis, distribution f, decay scaling
- eg if f has finite second moments, then  $u \in L_2$  almost surely iff  $(\gamma_\ell) \in \ell_2$
- $B_{11}^{s}$ -Besov priors:  $\xi_{\ell} \stackrel{iid}{\sim} Laplace(0,1)$  and  $\gamma_{\ell} = \ell^{\frac{1}{2} \frac{s}{d}}$ , s smoothness parameter

$$\pi(u)\propto \exp(-\|u\|_{B^s_{11}})$$
"

- M. Lassas, E. Saksman and S. Siltanen, *Discretization-invariant Bayesian inversion and Besov space priors, Inverse Problems and Imaging* 2009
- V. Kolehmainen, M. Lassas, K. Niinimaki and S. Siltanen, *Sparsity-promoting Bayesian inversion*, Inverse Prob 2012
- M. Dashti, S. Harris and A. Stuart, Besov priors for Bayesian inverse problems, Inverse Problems and Imaging, 2013

WNM - ROC under Besov regularity

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### Priors via random series expansions - Haar wavelets



- S. Agapiou, M. Burger, M. Dashti and T. Helin, *Sparsity-promoting and edge-preserving MAP estimators in nonparametric Bayesian inverse problems*, Inverse Problems, 2018.
- S. Agapiou, M. Dashti and T. Helin, *Rates of contraction of posterior distributions based on p-exponential priors*, to appear in Bernoulli.

## Outline

### 1 Motivation

### WNM - Minimax rates under Besov regularity

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### White noise model - Minimax estimation rates

• Observe solution to

$$dY_t^{(n)} = u(t)dt + \frac{1}{\sqrt{n}}dW_t, \quad t \in [0,1]$$

$$Y_0^{(n)} = 0$$
,  $W_t$  is a sBM

- $u \in L_2[0,1]$  unknown
- $P_u^{(n)}$  distribution of  $Y_t^{(n)}$
- Interested in small noise limit  $n \to \infty$

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### White noise model - Minimax estimation rates

• Minimax risk in  $L_2$ -loss over class  $\mathcal{F} \subset L_2[0,1]$ 

$$egin{aligned} R_{n}(\hat{u}, u) &= \min_{\hat{u}} \max_{u \in \mathcal{F}} \mathbb{E}_{P_{u}^{(n)}} \| \hat{u} - u \|_{L_{2}}^{2} \end{aligned}$$

- Minimax rate in L<sub>2</sub>-risk over  $\mathcal{F}$ : fastest rate of decay of above minimax risk, as  $n \to \infty$
- Linear minimax rate in  $L_2$ -risk over  $\mathcal{F}$ : restrict to linear estimators

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# WNM - Minimax estimation rates under Besov regularity

Theorem (Donoho + Johnstone '98)

In the WNM for  $\beta > \frac{1}{q}$  or  $\beta \ge 1$  for q = 1,

- Minimax rate in  $L_2$ -loss over  $B_{qq}^{\beta}$ 

$$m_n = n^{-\frac{\beta}{1+2\beta}}$$

- Linear minimax rate in  $L_2$ -loss over  $B_{qq}^{\beta}$ 

$$I_n = n^{-\frac{\beta - \gamma/2}{1 + 2\beta - \gamma}}$$

where 
$$\gamma = \frac{2}{q} - \frac{2}{q \vee 2} \ge 0$$
.

• For q < 2 (spatially inhomogeneous unknowns) linear estimators sub-optimal

e.g. 
$$q = 1$$
,  $I_n = n^{\frac{1-2\beta}{4\beta}}$ 

• Same result holds in Gaussian regression setting

D. Donoho and I. Johnstone, *Minimax estimation via wavelet shrinkage*, Annals of Statistics, 1998.

WNM - ROC under Besov regularity

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# NMR data denoising



Linear methods either oversmooth irregular part, or undersmooth regular part or both

I. Johnstone, Wavelets and the theory of non-parametric function estimation, Phil. tans. R. Soc. Lond. A, 1999.

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### *p*-exponential product measure

• 
$$\xi_\ell \stackrel{iid}{\sim} f_p$$
,  $f_p(x) = c_p e^{-\frac{|x|^p}{p}}$ ,  $p \in [1, 2]$ 

- $(\gamma_{\ell})$  decaying positive scalings
- Define *p*-exponential measure

$$\mu = \mathcal{L}\big((\gamma_\ell \xi_\ell)\big)$$

•  $\mu$  log-concave (unimodal, exponential moments, ...)

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## *p*-exponential product measure

- e.g.  $(\gamma_\ell) \in \ell_2$
- $\{\psi_\ell\}$  orthonormal basis in separable Hilbert-space X
- $\mu$  identified with measure in X via expansion

$$u(x) = \sum_{\ell=1}^{\infty} \gamma_\ell \xi_\ell \psi_\ell(x)$$

# Shift space

### Proposition (A., Dashti, Helin '20)

The space of admissible shifts of  $\mu$  is the Hilbert space

$$\mathcal{Q}(\mu) = \{h \in \mathbb{R}^{\infty} : \|h\|_{\mathcal{Q}} < \infty\},\$$

where

$$\|h\|_{\mathcal{Q}} = \Big(\sum_{\ell=1}^\infty rac{h_\ell^2}{\gamma_\ell^2}\Big)^{rac{1}{2}}.$$

For  $h \in \mathcal{Q}(\mu)$ 

$$\frac{d\mu(\cdot-h)}{d\mu}(u) = \lim_{N\to\infty} \prod_{\ell=1}^{N} \frac{f_p(u_\ell-h_\ell)}{f_p(u_\ell)} = \lim_{N\to\infty} e^{\frac{1}{p}\sum_{\ell=1}^{N} \left(|\frac{u_\ell}{\gamma_\ell}|^p - |\frac{u_\ell-h_\ell}{\gamma_\ell}|^p\right)}$$

- L. Shepp, *Distinguishing a sequence of random variables from a translate of itself*, Annals of Mathematical Statistics, 1965.
- S. Kakutani, On equivalence of infinite product measures, Annals of Mathematics, 1948.

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## Another important subspace

• Let 
$$\mathcal{Z}(\mu) = \{h \in \mathbb{R}^{\infty} : \|h\|_{\mathcal{Z}} < \infty\}$$
, where  
 $\|h\|_{\mathcal{Z}} = \Big(\sum_{\ell=1}^{\infty} |\frac{h_{\ell}}{\gamma_{\ell}}|^p\Big)^{\frac{1}{p}}$ 

•  $\mathcal{Z}$  Banach space

- $\mathcal{Z} \subset \mathcal{Q}$ , both null sets (e.g.  $\|(\gamma_{\ell}\xi_{\ell})\|_{\mathcal{Q}}^2 = \sum_{\ell=1}^{\infty} \xi_{\ell}^2$ )
- For Gaussian  $\mu$ :  $\mathcal{Z} = \mathcal{Q} = \mathcal{H}$ ,  $\mathcal{H}$  RKHS

• e.g.  $X = L_2[0, 1]$ ,  $Z \subset Q$  identified with spaces of functions with higher regularity

$$X = \operatorname{supp}(\mu) = \overline{\mathcal{Q}}^{\|\cdot\|_X} = \overline{\mathcal{Z}}^{\|\cdot\|_X}$$

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## Lower bound on probability of non-centered balls

### Theorem (A., Dashti, Helin '20)

For any  $h \in \mathcal{Z}$ 

$$\mu(\epsilon B_X + h) \geq e^{-\frac{1}{p} \|h\|_{\mathcal{Z}}^p} \mu(\epsilon B_X).$$



For proof:

- Use expression for  $\frac{d\mu(\cdot-h)}{d\mu}$
- Exploit symmetry and convexity (important that  $p \in [1,2]$ )

# Concentration function

• Define the concentration function for  $\mu$  a *p*-exponential measure at  $w \in X$ 

$$\phi_w(\epsilon) = \inf_{h \in \mathcal{Z}: \|h-w\|_X \le \epsilon} \frac{1}{p} \|h\|_{\mathcal{Z}}^p - \log \mu(\epsilon B_X)$$

- $\phi_0$  measures probability of  $\epsilon$ -balls around 0,  $\mu(\epsilon B_X) = e^{-\phi_0(\epsilon)}$
- Last theorem + approximation:

 $\phi_w$  controls probability of  $\epsilon$ -balls around  $w \in X$  from below

•  $\phi_w(\epsilon)$  blows up as  $\epsilon \to 0$ 

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## Talagrand's two level concentration inequality

### Lemma

There exists K > 0 depending only on p, s.t. for any  $\epsilon > 0$  and any M > 0 $\mu(\epsilon B_X + M^{\frac{p}{2}}B_Q + MB_Z) \ge 1 - \frac{1}{\mu(\epsilon B_X)}e^{-\frac{M^p}{K}}.$ 

M. Talagrand, *The supremum of some canonical processes*, American J. of Mathematics, 1994.

For Gaussian  $\mu$ , get Borell's concentration inequality

$$\mu(\epsilon B_X + MB_{\mathcal{H}}) \geq 1 - rac{1}{\mu(\epsilon B_X)}e^{-rac{M^2}{\kappa}}$$

C. Borell, The Brunn-Minkowski inequality in Gauss space, Inventiones Mathematicae, 1975.

# Rates of contraction

### Bayesian context with *p*-exponential priors

- Last two results allow to apply Ghosal and van der Vaart's ROC theory
  - Lower bound on prior probability around truth
  - Sieve set of bounded complexity, capturing most of prior mass
  - S. Ghosal and A. van der Vaart, *Convergence rates of posterior distributions for noniid observations*, Annals of Statistics, 2007.
  - A. van der Vaart and H. van Zanten, *Rates of contraction of posterior distributions based on Gaussian process priors*, Annals of Statistics, 2008.
- Control probability around truth using the concentration function
- Use  $\epsilon B_X + M^{\frac{p}{2}}B_Q + MB_Z$  as sieve set
  - Captures most of prior mass (Talagrand)
  - Complexity bound tricky because of having two large balls
  - If Q approximated well by  $\mathcal{Z}$  in X, can handle by embedding in  $2\epsilon B_X + cMB_Z$

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### 1 Motivation

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# WNM - Gaussian sequence model

• WNM or equivalently Gaussian sequence model

$$\gamma_\ell^{(n)} \stackrel{\textit{ind}}{\sim} \textit{N}(\textit{u}_\ell, 1/n), \quad \ell \in \mathbb{N}$$

- $u = (u_\ell) \in \ell_2$  unknown
- $P_u^{(n)}$  distribution of  $y^{(n)} = (y_\ell^{(n)})$
- $\bullet$  Interested in small noise limit,  $n \to \infty$

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# $\alpha$ -regular p-exponential priors

• Prior 
$$\Pi = \mathcal{L}((\gamma_{\ell}\xi_{\ell})), \ \xi_{\ell} \stackrel{iid}{\sim} f_{p}, \ p \in [1, 2]$$

• 
$$\gamma_{\ell} = \tau \ell^{-\frac{1}{2} - \alpha} \quad (\gamma_{kl} = \tau 2^{-(\frac{1}{2} + \alpha)k})$$

- $\tau > 0$  scaling parameter
- $\alpha > 0$  regularity parameter

### Lemma

For any  $q \ge 1$ , it holds  $\Pi(B^s_{qq}) = 1$  for all  $s < \alpha$  and  $\Pi(B^s_{qq}) = 0$  for all  $s \ge \alpha$ .

• For p=1,  $B_{11}^{lpha+1} ext{-Besov}$  prior

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## $\alpha$ -regular p-exponential priors

• Space of admissible shifts

$$\mathcal{Q} \coloneqq \mathcal{Q}_{\alpha} = \left\{ h \in \mathbb{R}^{\infty} : \|h\|_{\mathcal{Q}_{\alpha}} < \infty \right\}, \quad \|h\|_{\mathcal{Q}_{\alpha}} = \tau^{-1} \left( \sum_{\ell=1}^{\infty} \ell^{1+2\alpha} h_{\ell}^2 \right)^{\frac{1}{2}}$$

• Space determining mass-loss for noncentered balls

$$\mathcal{Z} \coloneqq \mathcal{Z}_{\alpha} = \{h \in \mathbb{R}^{\infty} : \|h\|_{\mathcal{Z}_{\alpha}} < \infty\}, \quad \|h\|_{\mathcal{Z}_{\alpha}} = \tau^{-1} \left(\sum_{\ell=1}^{\infty} \ell^{\frac{p}{2} + p\alpha} |h_{\ell}|^{p}\right)^{\frac{1}{p}}$$

• Identified with Besov spaces  $Q_{\alpha} = B_{22}^{\alpha + \frac{1}{2}}$  and  $Z_{\alpha} = B_{pp}^{\alpha + \frac{1}{p}}$ 

• Scaling parameter  $\tau$  appears in norms

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# WNM - Rates of contraction

Concentration function

$$\phi_w(\epsilon) = \inf_{h \in \mathcal{B}_{pp}^{\alpha + \frac{1}{p}} : \|h - w\|_{\ell_2} \le \epsilon} \frac{\tau^{-p}}{p} \|h\|_{\mathcal{B}_{pp}^{\alpha + \frac{1}{p}}}^p - \log \Pi(\epsilon B_{\ell_2})$$

### Theorem (A., Dashti, Helin '20)

Let  $u_0 \in \ell_2$  and let  $\epsilon_n \to 0$  such that

$$\phi_{u_0}(\epsilon_n) \leq n\epsilon_n^2.$$

Then as  $n \to \infty$ 

$$\exists_n(u : \|u-u_0\|_{\ell_2} > M\epsilon_n | y^{(n)}) \to 0,$$

in  $P_{u_0}^{(n)}$ -probability, for some M > 0.

# Estimating the concentration function

 $\bullet$  Centered small ball probabilities: for any  $\tau >$  0,  $\alpha >$  0 and  $\textit{p} \in$  [1,2]

$$-\log \Pi(\epsilon B_{\ell_2}) symp (\epsilon/ au)^{-rac{1}{lpha}}$$

- F. Aurzada, On the lower tail probabilities of some random sequences in  $\ell_p$ , J. Theoretical Probability, 2007.
- Decentering:

$$\inf_{\boldsymbol{p}\in \mathcal{B}_{pp}^{\alpha+\frac{1}{p}}:\|\boldsymbol{h}-\boldsymbol{u}_0\|_{\ell_2}\leq \epsilon}\frac{\tau^{-p}}{p}\|\boldsymbol{h}\|_{\mathcal{B}_{pp}^{\alpha+\frac{1}{p}}}^p$$

- $h_{1:L}$  truncation of  $u_0$  up to L,  $u_{1:L} \in B_{pp}^{lpha+rac{1}{p}}$
- Depending on regularity of  $u_0$ , for large enough L,  $\|h_{1:L} u_0\|_{\ell_2} \leq \epsilon$
- Depending on regularity of  $u_0$ , get bound on  $||h_{1:L}||_{B^{\alpha+\frac{1}{p}}_{oo}}$  hence also on infimum

## Rates under Sobolev regularity - no rescaling

### Theorem (A., Dashti, Helin '20)

Assume  $u_0 \in B_{22}^{\beta}$ . Consider  $\alpha$ -regular *p*-exponential priors  $p \in [1, 2]$ , with fixed  $\tau > 0$ . Then the posterior contracts at rate

$$\epsilon_{n} = \begin{cases} n^{-\frac{\beta}{1+2\beta+p(\alpha-\beta)}}, & \text{if } \alpha > \beta\\ n^{-\frac{\alpha}{1+2\alpha}}, & \text{if } \alpha \le \beta \end{cases}$$

• For 
$$\alpha = \beta$$
 minimax rate  $m_n = n^{-\frac{\beta}{1+2\beta}}$ 

- For  $\alpha < \beta$  rate independent of  $\textbf{\textit{p}}$ 
  - L. Zhao, *Bayesian aspects of some nonparametric problems*, Annals of Statistics, 2000.
  - I. Castillo and R. Nickl, *Nonparametric Bernstein von Mises theorems in Gaussian white noise*, Annals of Statistics, 2013.
- For  $\alpha > \beta$  the smaller p is the faster the rate

# Rates under Sobolev regularity - rescaling

### Theorem (A., Savva 21)

Assume  $u_0 \in B_{22}^{\beta}$ . Consider  $\alpha$ -regular *p*-exponential priors  $p \in [1, 2]$ , with  $\tau_n = \tau_n(\alpha, \beta, p)$  chosen optimally. Then the posterior contracts at rate

$$\epsilon_{n} := \begin{cases} n^{-\frac{\beta}{1+2\beta}}, & \text{if } \alpha > \beta - 1/p \\ n^{-\frac{\beta}{1+2\beta}} \log^{\frac{2-p}{2\beta p+p}} n, & \text{if } \alpha = \beta - 1/p \\ n^{-\frac{1+\alpha p}{2+p+2\alpha p}}, & \text{if } \alpha < \beta - 1/p. \end{cases}$$

• The smaller *p* is, the more undersmoothing the prior can be while still achieving the minimax rate

• Same picture for 
$$u_0 \in B^{eta}_{qq}, \ q>2$$

# Rates under spatially inhomogeneous truth - no rescaling

### Theorem (A., Dashti, Helin '20)

Assume  $u_0 \in B_{qq}^{\beta}$ , q < 2,  $\beta > 0 \lor (\frac{1}{q} - \frac{1}{2})$ . Consider  $\alpha$ -regular *p*-exponential priors  $p \in [1, 2]$ , with fixed  $\tau > 0$ . Then the best ROC  $\epsilon_n$  is achieved for

$$\alpha = \alpha(\beta, \boldsymbol{p}, \boldsymbol{q}) < \beta$$

and

- If  $p \leq q$ 

$$m_n \ll \epsilon_n \ll I_n.$$

As p grows towards q the best rate improves.

- If *q* < *p* < 2

$$m_n \ll \epsilon_n \ll I_n.$$

As p decreases towards q the best rate improves.

- If p = 2

$$\epsilon_n = I_n$$

 $(" \ll " = polynomially faster, I_n linear minimax rate)$ 

# Rates under spatially inhomogeneous truth - with rescaling

### Theorem (A., Dashti, Helin '20)

Assume  $u_0 \in B_{qq}^{\beta}$ , q < 2,  $\beta > \frac{1}{p} \lor \frac{1}{q}$ . Consider  $\alpha$ -regular *p*-exponential priors  $p \in [1, 2]$ , with  $\tau_n = \tau_n(\alpha, \beta, p, q)$  chosen optimally. Then the posterior contracts at rate  $\epsilon_n$  s.t.:

- For p = q,  $\alpha = \beta \frac{1}{p}$   $\epsilon_n = m_n$  for  $\tau_n = n^{-\frac{1}{p(1+2\beta)}}$ . - For p < q,  $\alpha = \beta - \frac{1}{p}$  $\epsilon_n = m_n \log^{\frac{q-p}{pq(1+2\beta)}} n$ .
- In all other cases

 $m_n \ll \epsilon_n$ .

- For p = 2 the best achievable rate is  $\epsilon_n = I_n$ , achieved for any  $\alpha \ge \beta - \frac{1}{q}$ .

e.g.  $u_0 \in B_{11}^1$  can use rescaled  $\alpha$ -regular Laplace prior, with  $\alpha > 0$  arbitrarily small, to get  $\epsilon_n = m_n n^{\delta}$  for  $\delta > 0$  arbitrarily small.

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# Difficulty for q < 2

Assume  $u_0 \in B_{qq}^{\beta}$ 

$$\sum_{\ell=1}^{\infty} \ell^{\beta q + \frac{q}{2} - 1} |u_{0,\ell}|^q < \infty$$

•  $h_{1:L}$  truncation of  $u_0$  up to L

$$\|h_{1:L} - u_0\|_{\ell_2}^2 = \sum_{\ell > L} u_{0,\ell}^2 = \sum_{\ell > L} \ell^{-2\beta - 1 + \frac{2}{q}} \ell^{2\beta + 1 - \frac{2}{q}} u_{0,\ell}^2$$

• If 
$$q=2$$
 can bound by  $\|u_0\|^2_{B^\beta_{22}}L^{-2\beta}$ 

- If q > 2 can use Hölder inequality  $(\frac{q}{2}, \frac{q}{q-2})$  to bound by  $||u_0||^2_{B^{\beta}_{aq}}L^{-2\beta}$
- If q < 2 cannot use Hölder inequality, forced to use crude bound

$$|u_{0,\ell}| \leq ||u_0||_{B^{\beta}_{qq}} \ell^{-\beta - \frac{1}{2} + \frac{1}{q}}$$

• Estimating  $\|h_{1:L}\|_{\mathcal{B}^{\alpha+\frac{1}{p}}_{pp}}$ : if  $p \leq q$  can use Hölder, if p > q need crude bound

# Difficulty for q < 2

$$\phi_{u_0}(\epsilon) \leq \tau^{-p} \epsilon^{-s} + (\epsilon/\tau)^{-\frac{1}{\alpha}}, \qquad s = s(\alpha, \beta, p, q) > 0$$

- Inf-term worse for larger  $\alpha$  (s increases with  $\alpha$ )
- $\bullet\,$  Centered small ball term better for larger  $\alpha$
- $\bullet\,$  To achieve minimax rate need centered term to dominate for  $\alpha\leq\beta$
- $\bullet$  Crude bound: inf-term dominates already for large enough  $\alpha < \beta$
- $\bullet$  Address by choosing small  $\alpha$  s.t. centered term dominates, inf-term better
  - Correct by sending  $\tau \rightarrow 0$  which improves centered term
  - Pay penalty  $\tau^{-p}$  in infimum term, hopefully less than choosing larger  $\alpha$
  - For p=2 penalty identical to increasing lpha, no improvement
  - For q improvement not enough to get minimax rate
  - For  $p \leq q$  enough improvement to get minimax rate (up to logs for p < q)

## Are Gaussian priors limited by the linear minimax rate?

In general minimax rate in  $L_2$ -risk benchmark for ROC in  $L_2$ 

- ROC  $\epsilon_n$  implies center of smallest ball containing half posterior mass, converges at same rate in probability
  - S. Ghosal, J. Ghosh and A. van der Vaart, *Convergence rates of posterior distributions*, Annals of Statistics, 2000.
- Lower bounds on minimax rates in L<sub>2</sub>-risk proved by showing lower bounds in probability (stronger statement due to Markov inequality)

## Are Gaussian priors limited by the linear minimax rate?

For Gaussian priors, hence Gaussian posteriors

- Posterior mean = MAP estimator, linear estimator
- ROC  $\epsilon_n$  implies posterior mean converges at same rate in probability
- Lower bounds on linear minimax rate derived in  $L_2$ -risk directly
- Need to show that posterior mean converges at same rate also in  $L_2$ -risk

(S. Wang)

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## NMR data

- Nuclear Magnetic Resonance data, available in WaveLab 850
- Signal expanded in Symlet 6 orthonormal wavelet basis  $\{\psi_{kl}\}$  truncated at k = 9



# Bayesian Denoising of NMR data

• Model wavelet coefficients as

$$y_{kl} = u_{kl} + rac{1}{\sqrt{\delta}} z_{kl}, \quad z_{kl} \stackrel{iid}{\sim} N(0,1)$$

• Rescaled  $\alpha$ -regular *p*-exponential prior on unknown  $u = (u_{kl})$ , with p = 1 or 2

$$\mu_{kl}= au 2^{-(rac{1}{2}+lpha)k}\xi_{kl}, \quad \xi_{kl}\stackrel{iid}{\sim} f_p, \ p=1 ext{ or } 2^{-(rac{1}{2}+lpha)k}\xi_{kl}$$

• Hyperprior on prior-rescaling au:  $au^{-2} \sim \text{Gamma}(a_1, b_1)$ 

- Hyperprior on noise-precision  $\delta$ :  $\delta \sim \text{Gamma}(a_2, b_2)$
- $a_1, a_2, b_1, b_2$  chosen so that hyperpriors non-informative for  $\tau, \delta$

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# Bayesian Denoising of NMR data - Gaussian prior

- Conditional conjugacy
  - $u_{kl}|y_{kl}, \tau, \delta \sim N(m_{kl}, c_{kl})$
  - $\tau^{-2}|u, y \sim \operatorname{Gamma}(a'_1, b'_1(u))$
  - $\delta | u, y \sim \text{Gamma}(a'_2, b'_2(u, y))$
- Can use simple Gibbs Sampler to sample posterior
- Normally in high-dim  $\tau$ -chain mixes poorly (u and  $\tau$  a-priori strongly dependent)  $\rightarrow$  use non-centered parametrization  $u = \tau v$ , and work with v instead of u
- S. Agapiou, J. Bardsley, O. Papaspiliopoulos, A. Stuart *Analysis of the Gibbs Sampler for Hierarchical Inverse Problems*, SIAM/ASA Journal on UQ, 2014.

# Bayesian Denoising of NMR data - Laplace prior

- No conditional conjugacy (only for  $\delta | u, y$ )
- Need to use Metropolis within Gibbs
- pCN dimension-robust for Gaussian priors
- Again  $u, \tau$  a-priori strongly dependent
- $\bullet$  Use non-centered pCN within Gibbs
  - Write  $u = T(\zeta, \tau)$  such that  $\zeta, \tau$  a-priori independent and  $\zeta$  is Gaussian WN
  - Sample iteratively  $\zeta|y, \tau$  (pCN) and  $\tau|y, \zeta$  (independence sampler)
- V. Chen, M. Dunlop, O. Papaspiliopoulos, A. Stuart *Dimension-Robust MCMC in Bayesian Inverse Problems*, arXiv:1803.03344.

# NMR data - Gauss vs Laplace priors



### NMR data - Gauss vs Laplace priors - $\tau$ -chains



## NMR data - Gauss vs Laplace priors - $\delta$ -chains



# Outline

### **Motivation**

WNM - Minimax rates under Besov regularity

### *p*-exponential measures

WNM - ROC under Besov regularity

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# Summary and open questions

- Laplace priors appear to outperform Gaussian priors over Besov regularity
- Sharpness of rates
- For benefit to be realized need better algorithms
- Adaptation over Besov spaces (with A. Savva)
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### THANK YOU!

# http://www.sergiosagapiou.com/

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