

Adaptive rates of contraction for spatially inhomogeneous unknowns




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Joint work with

-  S. Agapiou, M. Dashti and T. Helin, *Rates of contraction of posterior distributions based on p -exponential priors*, Bernoulli, 2021.
-  S. Agapiou and S. Wang, *Laplace priors and spatial inhomogeneity in Bayesian inverse problems*, arXiv:2112.05679.
-  S. Agapiou and A. Savva, *Adaptive rates of contraction based on p -exponential priors*, in preparation.

Outline

- 1 Motivation
- 2 WNM - Minimax rates under Besov regularity
- 3 p -exponential measures
- 4 WNM - ROC under Besov regularity
- 5 Numerics
- 6 Conclusion

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Multiscale features in images



Wavelet expansions

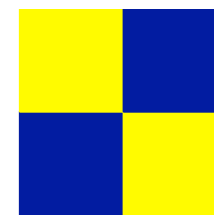
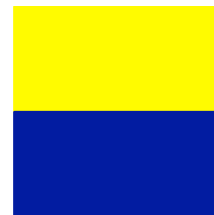
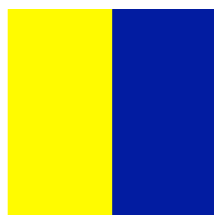
- $\{\psi_\ell\}_{\ell=1}^\infty$ orthonormal basis for L_2

$$u(x) = \sum_{\ell=1}^{\infty} u_\ell \psi_\ell(x), \quad u_\ell = \langle u, \psi_\ell \rangle.$$

- e.g. $\{\psi_\ell\}$ is the Fourier basis
- For functions with multiscale features, better use **wavelet** bases $\{\psi_{kl}\}$

$$u(x) = \sum_{k=1}^{\infty} \sum_{l=1}^{2^k} u_{kl} \psi_{kl}(x), \quad u_{kl} = \langle u, \psi_{kl} \rangle.$$

e.g. 2D Haar



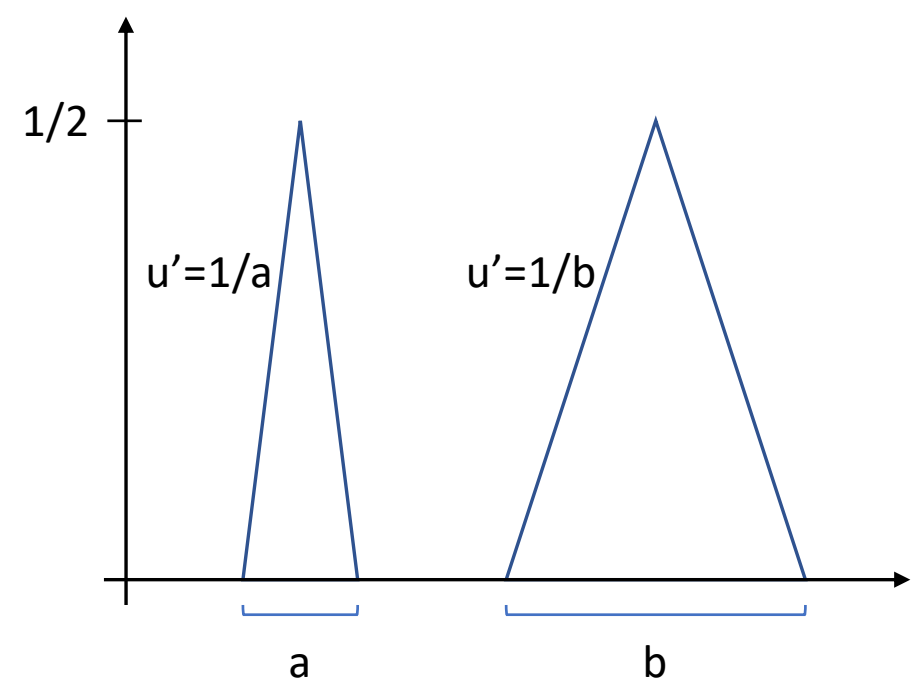
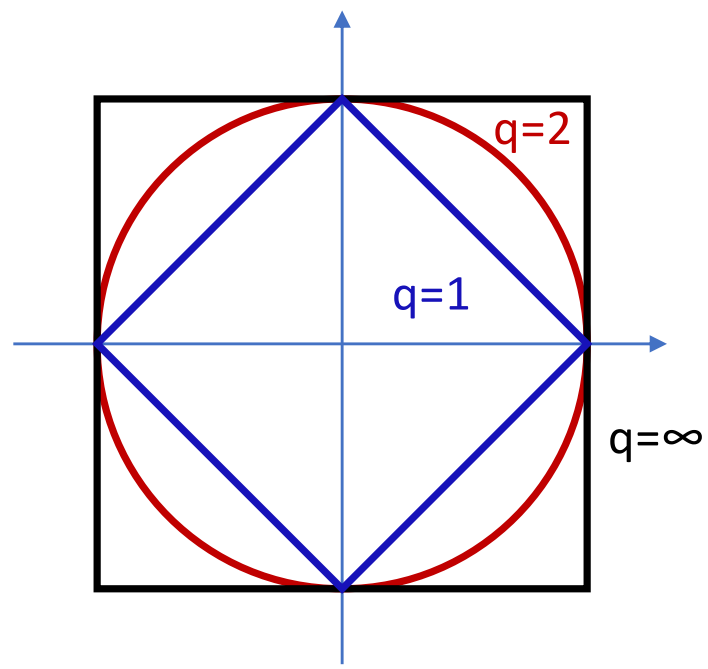
Besov Spaces

- Functions identified with expansion coefficients $(u_\ell) \in \ell_2$ or $(u_{kl}) \in \ell_2$
- Besov space of **smoothness** $s \in \mathbb{R}$, with **integrability** parameter $q \geq 1$

$$B_{qq}^s = \left\{ u \in \mathbb{R}^\infty : \sum_{\ell=1}^{\infty} \ell^{q(\frac{s}{d} + \frac{1}{2}) - 1} |u_\ell|^q < \infty \right\}, \quad \|u\|_{B_{qq}^s} = \left(\sum_{\ell=1}^{\infty} \ell^{q(\frac{s}{d} + \frac{1}{2}) - 1} |u_\ell|^q \right)^{\frac{1}{q}}.$$

- $q = 2$: $B_{22}^s = H^s$, Sobolev Hilbert spaces
- $q = \infty$, $s \notin \mathbb{N}$: $B_{\infty\infty}^s = C^s$, Hölder spaces
- Smaller q associated with **sparsity** and **spatial inhomogeneity**

Besov Spaces - Intuition



$$\|u'\|_{L_1} = 2, \quad \|u'\|_{L_2} = \sqrt{\frac{1}{a} + \frac{1}{b}}, \quad \|u'\|_{L_\infty} = \frac{1}{a}$$

📄 I.M. Johnstone, *Gaussian estimation: sequence and wavelet models*, draft book.

Function-space priors via random series expansions

$$u(x) = \sum_{\ell=1}^{\infty} u_{\ell} \psi_{\ell}(x)$$

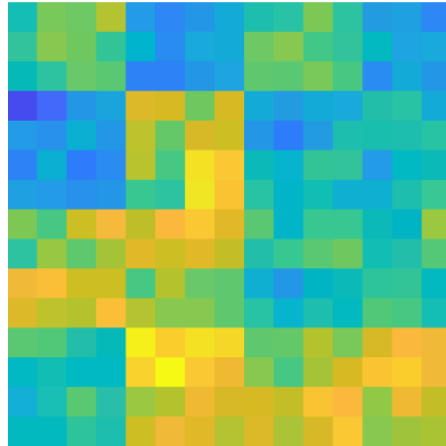
- **Randomize** coefficients: $u_{\ell} = \gamma_{\ell} \xi_{\ell}$ where $\xi_{\ell} \stackrel{iid}{\sim} f$, $\gamma_{\ell} > 0$ decaying scalings
- Choice of wavelet basis, distribution f , decay scaling
- eg if f has finite second moments, then $u \in L_2$ almost surely iff $(\gamma_{\ell}) \in \ell_2$
- **B_{11}^s -Besov priors:** $\xi_{\ell} \stackrel{iid}{\sim} \text{Laplace}(0, 1)$ and $\gamma_{\ell} = \ell^{\frac{1}{2} - \frac{s}{d}}$, s smoothness parameter

$$" \pi(u) \propto \exp(-\|u\|_{B_{11}^s}) "$$

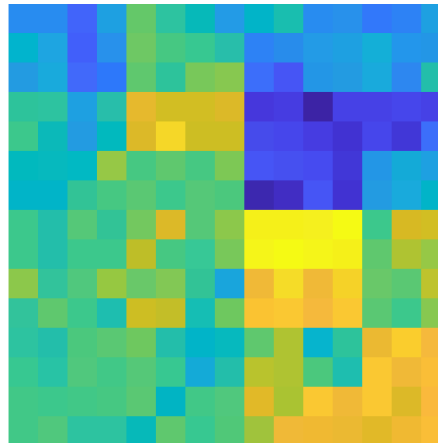


M. Lassas, E. Saksman and S. Siltanen, *Discretization-invariant Bayesian inversion and Besov space priors*, *Inverse Problems and Imaging* 2009

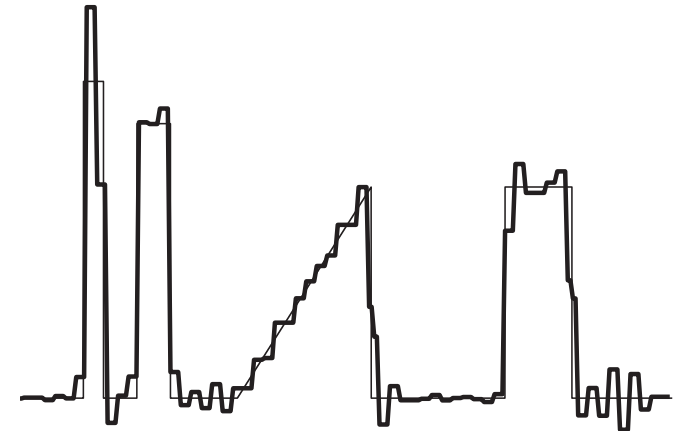
Priors via random series expansions - Haar wavelets



Gaussian



Laplace (B_{11}^s)



Kolehmainen et al. 2012

- 📄 S. Agapiou, M. Burger, M. Dashti and T. Helin, *Sparsity-promoting and edge-preserving MAP estimators in nonparametric Bayesian inverse problems*, Inverse Problems, 2018.
- 📄 V. Kolehmainen, M. Lassas, K. Niinimäki and S. Siltanen, *Sparsity-promoting Bayesian inversion*, Inverse Prob 2012

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White noise model - Minimax estimation rates

- Observe solution to

$$dY_t^n = u(t)dt + \frac{1}{\sqrt{n}}dW_t, \quad t \in [0, 1]$$

$Y_0^n = 0$, W_t is a sBM

- $u \in L_2[0, 1]$ **unknown**, work in ℓ_2
- P_u^n distribution of Y^n
- Interested in small noise limit $n \rightarrow \infty$

White noise model - Minimax estimation rates

- Minimax risk in ℓ_2 -loss over class $\mathcal{F} \subset \ell_2$

$$R_n(\hat{u}, u) = \min_{\hat{u}} \max_{u \in \mathcal{F}} \mathbb{E}_{P_u^n} \|\hat{u} - u\|_2^2$$

- **Minimax rate in ℓ_2 -risk** over \mathcal{F} : fastest rate of decay of minimax risk, as $n \rightarrow \infty$
- **Linear** minimax rate in ℓ_2 -risk over \mathcal{F} : restrict to linear estimators

WNM - Minimax estimation rates under Besov regularity

Theorem (Donoho + Johnstone '98)

In the WNM for $\beta > \frac{1}{q}$ or $\beta \geq 1$ for $q = 1$,

- **Minimax rate** in ℓ_2 -loss over B_{qq}^β

$$m_n = n^{-\frac{\beta}{1+2\beta}}$$

- Linear minimax rate in ℓ_2 -loss over B_{qq}^β

$$l_n = n^{-\frac{\beta-\gamma/2}{1+2\beta-\gamma}},$$

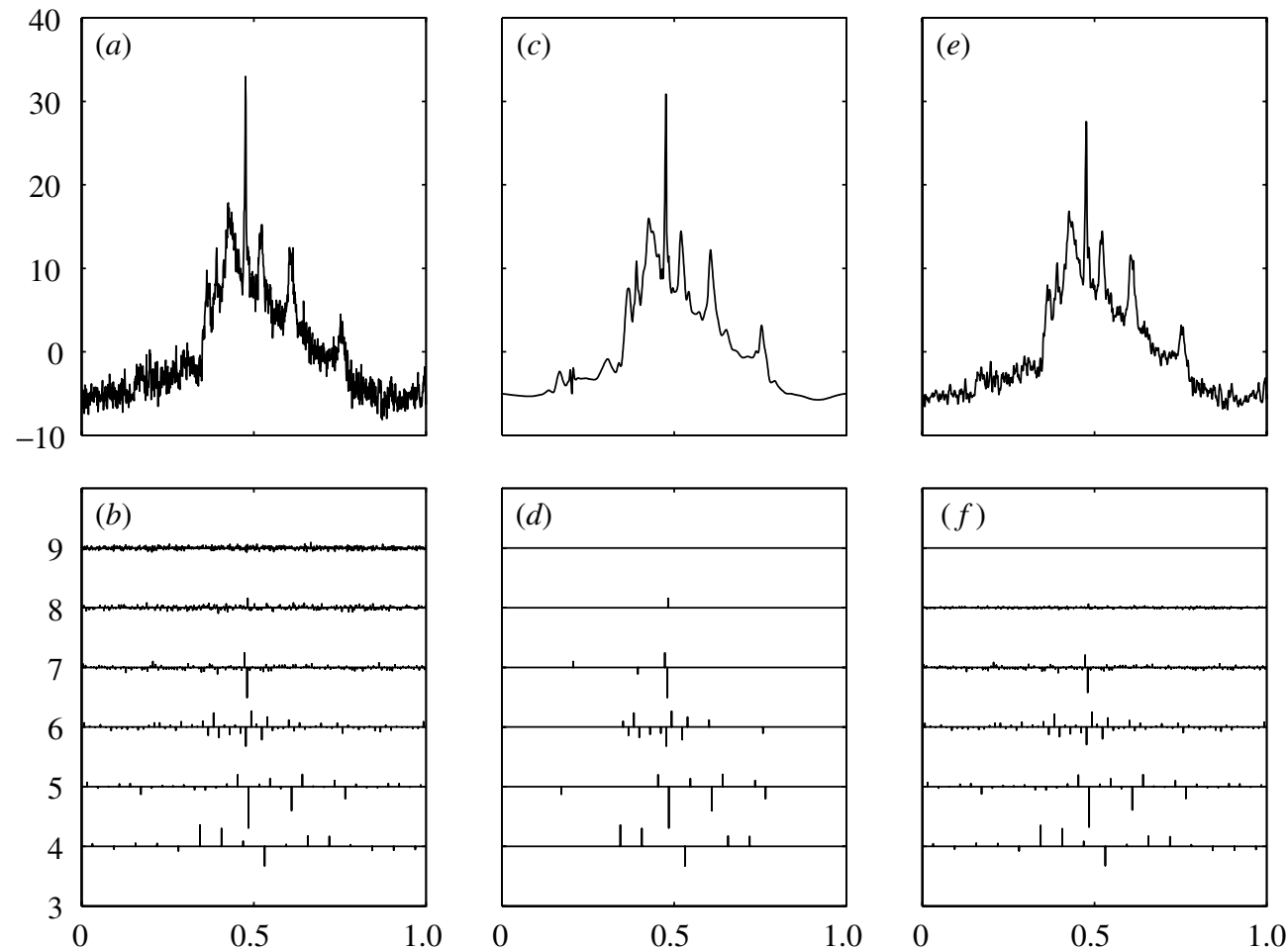
where $\gamma = \frac{2}{q} - \frac{2}{q\sqrt{2}} \geq 0$.

- For $q < 2$ (**spatially inhomogeneous unknowns**) linear estimators **sub-optimal**
- Same result holds in Gaussian regression setting



D. Donoho and I. Johnstone, *Minimax estimation via wavelet shrinkage*, Annals of Statistics, 1998.

NMR data denoising



Linear methods either **oversmooth** irregular part, or **undersmooth** regular part or **both**

📄 I. Johnstone, *Wavelets and the theory of non-parametric function estimation*, Phil. trans. R. Soc. Lond. A, 1999.

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p-exponential product measure

- $\xi_\ell \stackrel{iid}{\sim} f_p, \quad f_p(x) = c_p e^{-\frac{|x|^p}{p}}, \quad p \in [1, 2]$

- (γ_ℓ) decaying positive scalings

- Define *p*-exponential measure

$$\Pi = \mathcal{L}((\gamma_\ell \xi_\ell))$$

- Π log-concave (unimodal, exponential moments, ...)

Frequentist performance of posterior

- Prior Π on $u \in \ell_2$
- Posterior $\Pi(\cdot | Y^n)$ on u
- Frequentist assumption: observations Y^n in WNM generated from fixed $u_0 \in \ell_2$
- ϵ_n is a **posterior contraction rate** at u_0 , if $\exists M > 0$ such that as $n \rightarrow \infty$

$$\Pi(u : \|u - u_0\|_2 \geq M\epsilon_n | Y^n) \rightarrow 0$$

in $P_{u_0}^n$ -probability

- Do Gaussian priors perform better for Sobolev truths?
- Do Laplace priors perform better for spatially inhomogeneous truths?

Rates of contraction - General Theory

General contraction theory \rightarrow rate ϵ_n depends on

- Prior putting a certain minimum mass on small ℓ_2 -balls around u_0
- Existence of sieve sets such that:
 - capture the bulk of prior's mass
 - their elements can be tested against u_0 with good enough type I & type II errors



S. Ghosal and A. van der Vaart, *Convergence rates of posterior distributions for noniid observations*, Annals of Statistics, 2007.



S. Ghosal and A. van der Vaart, *Fundamentals of nonparametric Bayesian inference*, Cambridge Series in Statistical and Probabilistic Mathematics, 2017.

Shift space

Proposition (A., Dashti, Helin '21)

The space of admissible shifts of Π is the Hilbert space


$$\mathcal{Q} = \{h \in \mathbb{R}^\infty : \|h\|_{\mathcal{Q}} < \infty\},$$

where

$$\|h\|_{\mathcal{Q}} = \left(\sum_{\ell=1}^{\infty} \frac{h_\ell^2}{\gamma_\ell^2} \right)^{\frac{1}{2}}.$$

For $h \in \mathcal{Q}$

$$\frac{d\Pi(\cdot - h)}{d\Pi}(u) = \lim_{N \rightarrow \infty} \prod_{\ell=1}^N \frac{f_p(u_\ell - h_\ell)}{f_p(u_\ell)} = \lim_{N \rightarrow \infty} e^{\frac{1}{p} \sum_{\ell=1}^N \left(| \frac{u_\ell}{\gamma_\ell} |^p - | \frac{u_\ell - h_\ell}{\gamma_\ell} |^p \right)}.$$

 L. Shepp, *Distinguishing a sequence of random variables from a translate of itself*, Annals of Mathematical Statistics, 1965.

 S. Kakutani, *On equivalence of infinite product measures*, Annals of Mathematics, 1948.

Another important subspace

- Let $\mathcal{Z} = \{h \in \mathbb{R}^\infty : \|h\|_{\mathcal{Z}} < \infty\}$, where

$$\|h\|_{\mathcal{Z}} = \left(\sum_{l=1}^{\infty} \left| \frac{h_l}{\gamma_l} \right|^p \right)^{\frac{1}{p}}$$

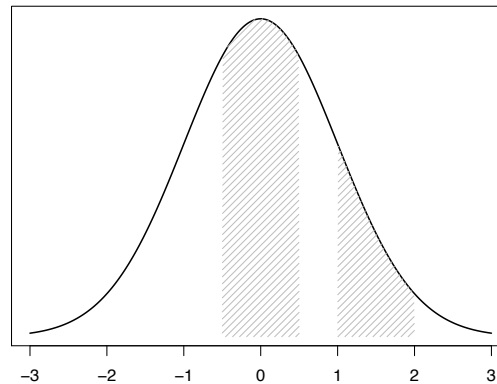
- \mathcal{Z} Banach space
- $\mathcal{Z} \subset \mathcal{Q}$, both null sets (e.g. $\|(\gamma_l \xi_l)\|_{\mathcal{Q}}^2 = \sum_{l=1}^{\infty} \xi_l^2$)
- For Gaussian Π : $\mathcal{Z} = \mathcal{Q} = \mathcal{H}$, \mathcal{H} RKHS

Lower bound on probability of non-centered balls

Theorem (A., Dashti, Helin '21)

For any $h \in \mathcal{Z}$

$$\Pi(\epsilon B_{\ell_2} + h) \geq e^{-\frac{1}{p} \|h\|_{\mathcal{Z}}^p} \Pi(\epsilon B_{\ell_2}).$$



For proof:

- Use expression for $\frac{d\Pi(\cdot - h)}{d\Pi}$
- Exploit symmetry and convexity (important that $p \in [1, 2]$)

Concentration function

- Define the **concentration function** for Π a p -exponential measure at $w \in \ell_2$

$$\phi_w(\epsilon) = \inf_{h \in \mathcal{Z}: \|h-w\|_2 \leq \epsilon} \frac{1}{p} \|h\|_{\mathcal{Z}}^p - \log \Pi(\epsilon B_{\ell_2})$$

- ϕ_0 measures probability of ϵ -balls around 0, $\Pi(\epsilon B_{\ell_2}) = e^{-\phi_0(\epsilon)}$
- Last theorem + approximation:

ϕ_w controls probability of ϵ -balls around $w \in \ell_2$ from below


- Note that $\phi_w(\epsilon) \rightarrow 0$, as $\epsilon \rightarrow 0$

Talagrand's two level concentration inequality

Lemma


There exists $K > 0$ depending only on p , s.t. for any $\epsilon > 0$ and any $M > 0$

$$\Pi(\epsilon B_{\ell_2} + M^{\frac{p}{2}} B_Q + MB_Z) \geq 1 - \frac{1}{\Pi(\epsilon B_{\ell_2})} e^{-\frac{MP}{K}}.$$

 M. Talagrand, *The supremum of some canonical processes*, American J. of Mathematics, 1994.

For Gaussian Π , get Borell's concentration inequality

$$\Pi(\epsilon B_{\ell_2} + MB_{\mathcal{H}}) \geq 1 - \frac{1}{\Pi(\epsilon B_{\ell_2})} e^{-\frac{M^2}{K}}$$

 C. Borell, *The Brunn-Minkowski inequality in Gauss space*, Inventiones Mathematicae, 1975.

Rates of contraction

- Recall Ghosal and van der Vaart's ROC theory
 - Lower bound on prior probability around truth
 - Sieve set of bounded complexity, capturing most of prior mass
- Control probability around truth using the concentration function
- Use $\epsilon B_{\ell_2} + M^{\frac{p}{2}} B_Q + MB_Z$ as sieve set
 - Captures most of prior mass (Talagrand)
 - Concentration function turns out to control complexity as well

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α -regular τ -scaled p -exponential priors

- Prior $\Pi = \mathcal{L}((\gamma_\ell \xi_\ell))$, $\xi_\ell \stackrel{iid}{\sim} f_p$, $p \in [1, 2]$
- $\gamma_\ell = \tau \ell^{-\frac{1}{2}-\alpha}$ ($\gamma_{kl} = \tau 2^{-(\frac{1}{2}+\alpha)k}$)
- $\tau > 0$ **scaling** parameter
- $\alpha > 0$ **regularity** parameter

Lemma

For any $q \geq 1$, it holds $\Pi(B_{qq}^s) = 1$ for all $s < \alpha$ and $\Pi(B_{qq}^s) = 0$ for all $s \geq \alpha$.

α -regular τ -scaled p -exponential priors

- Space of admissible shifts $\mathcal{Q} := \mathcal{Q}_\alpha$

$$\|h\|_{\mathcal{Q}_\alpha} = \tau^{-1} \left(\sum_{\ell=1}^{\infty} \ell^{1+2\alpha} h_\ell^2 \right)^{\frac{1}{2}}$$

- Space determining mass-loss for noncentered ball $\mathcal{Z} := \mathcal{Z}_\alpha$

$$\|h\|_{\mathcal{Z}_\alpha} = \tau^{-1} \left(\sum_{\ell=1}^{\infty} \ell^{\frac{p}{2} + p\alpha} |h_\ell|^p \right)^{\frac{1}{p}}$$

- Identified with Besov spaces $\mathcal{Q}_\alpha = B_{22}^{\alpha + \frac{1}{2}}$ and $\mathcal{Z}_\alpha = B_{pp}^{\alpha + \frac{1}{p}}$


- Concentration function

$$\phi_w(\epsilon) = \inf_{h \in B_{pp}^{\alpha + \frac{1}{p}} : \|h - w\|_{\ell_2} \leq \epsilon} \frac{\tau^{-p}}{p} \|h\|_{B_{pp}^{\alpha + \frac{1}{p}}}^p - \log \Pi(\epsilon B_{\ell_2})$$

Estimating the concentration function

- Centered small ball probabilities: for any $\tau > 0$, $\alpha > 0$ and $p \in [1, 2]$

$$-\log \Pi(\epsilon B_{\ell_2}) \asymp (\epsilon/\tau)^{-\frac{1}{\alpha}}$$

 F. Aurzada, *On the lower tail probabilities of some random sequences in ℓ_p* , J. Theoretical Probability, 2007.

- Decentering:

$$\inf_{h \in B_{pp}^{\alpha + \frac{1}{p}} : \|h - u_0\|_{\ell_2} \leq \epsilon} \frac{\tau^{-p}}{p} \|h\|_{B_{pp}^{\alpha + \frac{1}{p}}}^p$$

- $h_{1:L}$ truncation of u_0 up to L , $u_{1:L} \in B_{pp}^{\alpha + \frac{1}{p}}$
- Depending on regularity of u_0 , for large enough L , $\|h_{1:L} - u_0\|_{\ell_2} \leq \epsilon$
- Depending on regularity of u_0 , get bound on $\|h_{1:L}\|_{B_{pp}^{\alpha + \frac{1}{p}}}$ hence also on infimum

Rates under Sobolev regularity

Theorem (A., Dashti, Helin '21)

Assume $u_0 \in B_{22}^\beta$ and consider an α -regular τ -scaled p -exponential prior $p \in [1, 2]$.
Then if either

- $\alpha = \beta$ with $\tau > 0$ fixed, or
- $\alpha > \beta - \frac{1}{p}$ and $\tau = \tau(n; \alpha, \beta, p)$ chosen optimally

the posterior contracts at the **minimax** rate $m_n = n^{-\frac{\beta}{1+2\beta}}$.

Rates under Sobolev regularity - adaptation

Same rates with data driven choice of α or τ (no a priori knowledge of β required)

- Hierarchical Bayes on smoothness α , e.g. using exponential hyper-prior
- Hierarchical Bayes on scaling τ , e.g. using inverse gamma hyper-priors
- Empirical Bayes, estimate α or τ using the maximum marginal likelihood estimator
- In preparation, with A. Savva



B. T. Szabó, A. W. van der Vaart, and J. H. van Zanten, *Empirical Bayes scaling of Gaussian priors in the white noise model*, Electronic Journal of Statistics, 2013.



B. T. Knapik, B. Szabó, A. W. van der Vaart, and J. van Zanten, *Bayes procedures for adaptive inference in inverse problems for the white noise model*, Probability Theory and Related Fields, 2016.



J. Rousseau, B. Szabó, *Asymptotic behaviour of the empirical Bayes posteriors associated to maximum marginal likelihood estimator*, The Annals of Statistics, 2017.

Rates under spatially inhomogeneous truth

Theorem (A., Dashti, Helin '21)

Assume $u_0 \in B_{qq}^\beta$, $q < 2$, $\beta > \frac{1}{p} \vee \frac{1}{q}$. Consider an α -regular τ -scaled p -exponential prior $p \in [1, 2]$, with $\tau_n = \tau_n(\alpha, \beta, p, q)$ chosen optimally. Then the posterior contracts at rate ϵ_n s.t.:

- For $p = q$, $\alpha = \beta - \frac{1}{p}$

$$\epsilon_n = m_n.$$

- For $p < q$, $\alpha = \beta - \frac{1}{p}$

$$\epsilon_n = m_n \log^{\frac{q-p}{pq(1+2\beta)}} n.$$

- In all other cases

$$\epsilon_n \gg m_n.$$

- For $p = 2$ the best achievable rate is $\epsilon_n = l_n \gg m_n$ (l_n linear minimax).

Appropriately tuned Laplace priors achieve **minimax** rate for $q < 2$ (up to logs if $q > 1$)

Suboptimality of Gaussian priors for spatial inhomogeneity

Theorem (A. and Wang '21)

Assume $\beta > 1/q$, $1 \leq q < 2$ or $\beta = q = 1$, and let $\delta_n \downarrow 0$ as $n \rightarrow \infty$.

Let $(\Pi_n : n \in \mathbb{N})$ be mean-zero Gaussian priors supported on L_2 , such that for all $\eta > 0$

$$\sup_{u_0: \|u_0\|_{B_{qq}^\beta} \leq 1} P_{u_0}^n(\Pi_n(u : \|u - u_0\|_2 \geq \delta_n | Y_n) \geq \eta) \xrightarrow{n \rightarrow \infty} 0.$$

Then there exists some constant $c > 0$ such that

$$\delta_n \geq cl_n, \quad n \in \mathbb{N}.$$

- Uniform statement on contraction rate required to link to minimax
- Tuned Laplace priors satisfy uniform contraction with $\delta_n = m_n \ll l_n!$

Gaussian vs Laplace priors

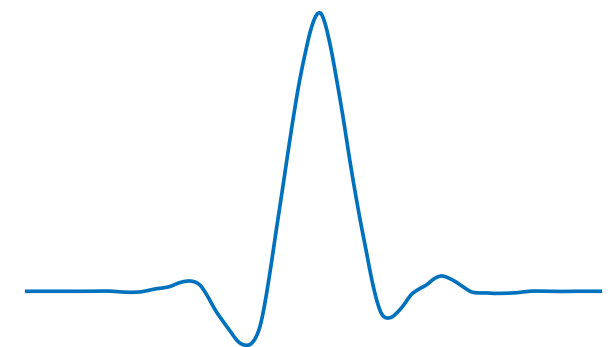
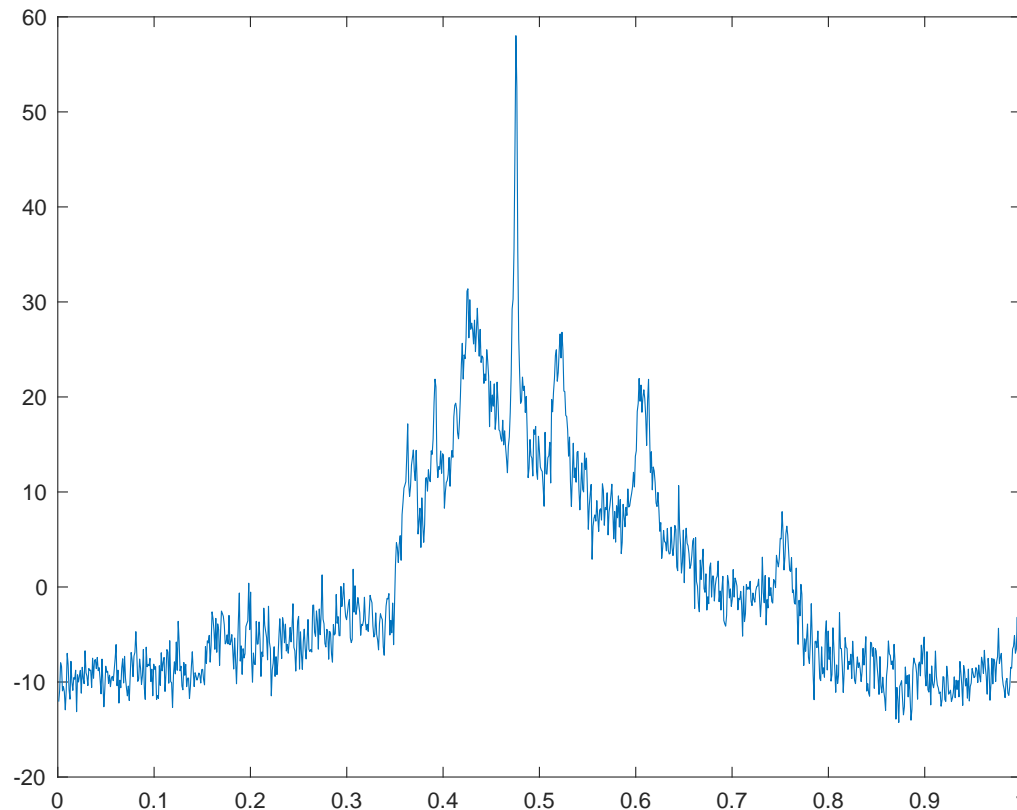
- For Sobolev truths Gaussian and Laplace priors have similar performance
- For spatially inhomogeneous truths, tuned Laplace priors **outperform** Gaussians
- Tuning, smoothness and scaling simultaneously, can be performed **adaptively** using Hierarchical or Empirical Bayes approach (in preparation with A. Savva)

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NMR data

- Nuclear Magnetic Resonance data, available in WaveLab 850
- Signal expanded in Symlet 6 orthonormal wavelet basis $\{\psi_{kl}\}$ truncated at $k = 9$



Bayesian Denoising of NMR data

- Model wavelet coefficients as

$$y_{kl} = u_{kl} + \frac{1}{\sqrt{\delta}} z_{kl}, \quad z_{kl} \stackrel{iid}{\sim} N(0, 1)$$

- Rescaled α -regular p -exponential prior on unknown $u = (u_{kl})$, with $p = 1$ or 2

$$u_{kl} = \tau 2^{-(\frac{1}{2} + \alpha)k} \xi_{kl}, \quad \xi_{kl} \stackrel{iid}{\sim} f_p, \quad p = 1 \text{ or } 2$$

- Hyperprior on prior-rescaling τ : $\tau^{-2} \sim \text{Gamma}(a_1, b_1)$
- Hyperprior on noise-precision δ : $\delta \sim \text{Gamma}(a_2, b_2)$
- a_1, a_2, b_1, b_2 chosen so that hyperpriors non-informative for τ, δ

Bayesian Denoising of NMR data - Gaussian prior

- Conditional conjugacy
 - $u_{kl} | y_{kl}, \tau, \delta \sim N(m_{kl}, c_{kl})$
 - $\tau^{-2} | u, y \sim \text{Gamma}(a'_1, b'_1(u))$
 - $\delta | u, y \sim \text{Gamma}(a'_2, b'_2(u, y))$
- Can use simple Gibbs Sampler to sample posterior
- Normally in high-dim τ -chain mixes poorly (u and τ a-priori strongly dependent)
 - use non-centered parametrization $u = \tau v$, and work with v instead of u



S. Agapiou, J. Bardsley, O. Papaspiliopoulos, A. Stuart *Analysis of the Gibbs Sampler for Hierarchical Inverse Problems*, SIAM/ASA Journal on UQ, 2014.

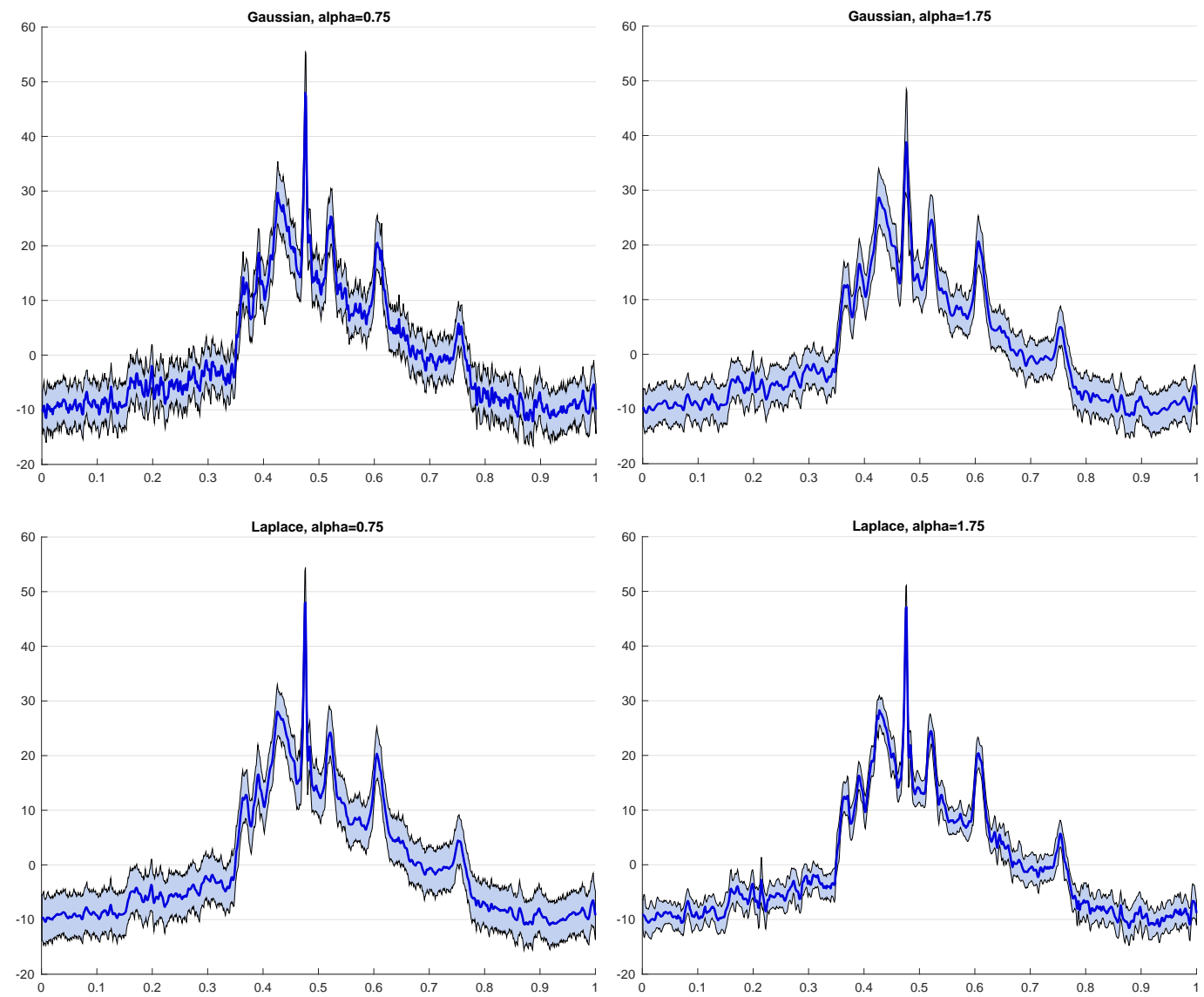
Bayesian Denoising of NMR data - Laplace prior

- No conditional conjugacy (only for $\delta|u, y$)
- Need to use Metropolis within Gibbs
- pCN dimension-robust for Gaussian priors
- Again u, τ a-priori strongly dependent
- Use **non-centered pCN within Gibbs**
 - Write $u = T(\zeta, \tau)$ such that ζ, τ a-priori independent and ζ is Gaussian WN
 - Sample iteratively $\zeta|y, \tau$ (pCN) and $\tau|y, \zeta$ (independence sampler)

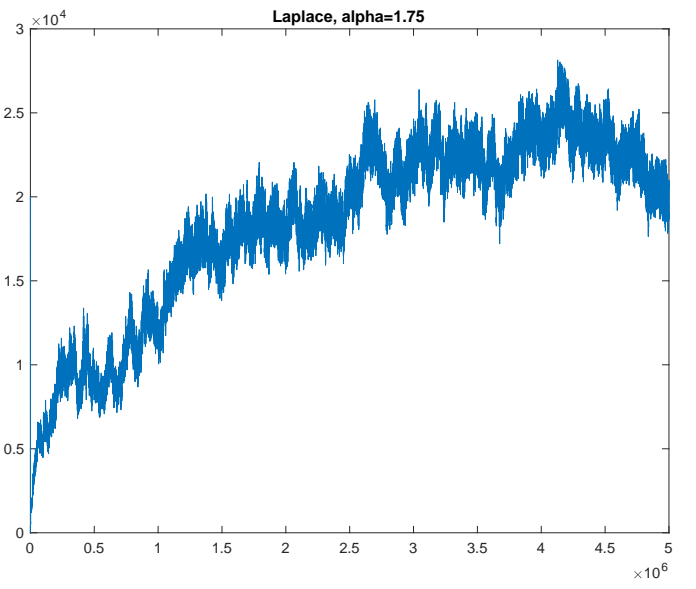
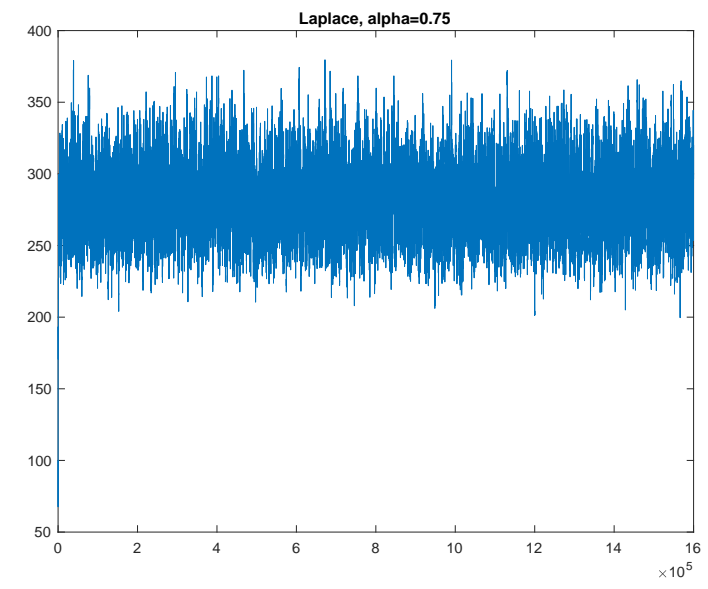
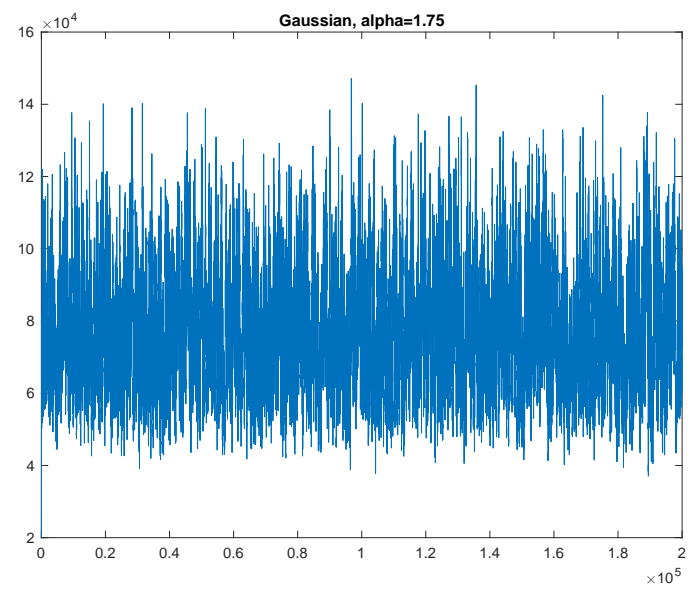
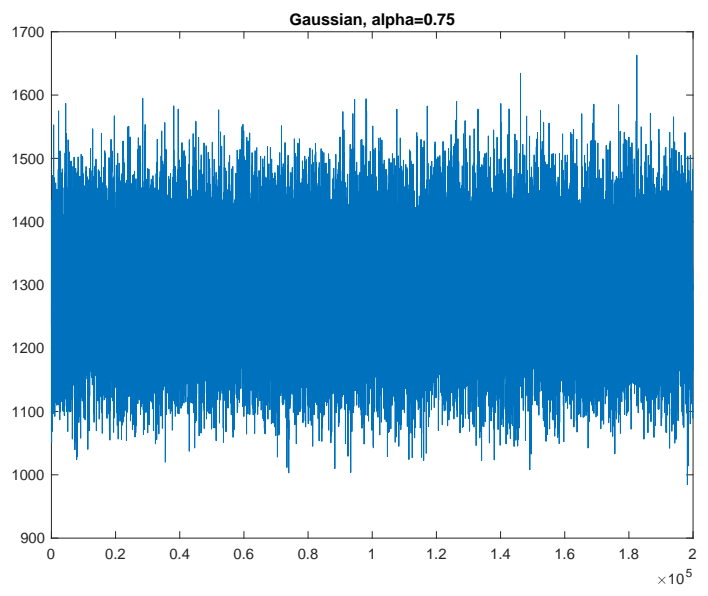


V. Chen, M. Dunlop, O. Papaspiliopoulos, A. Stuart *Dimension-Robust MCMC in Bayesian Inverse Problems*, arXiv:1803.03344.

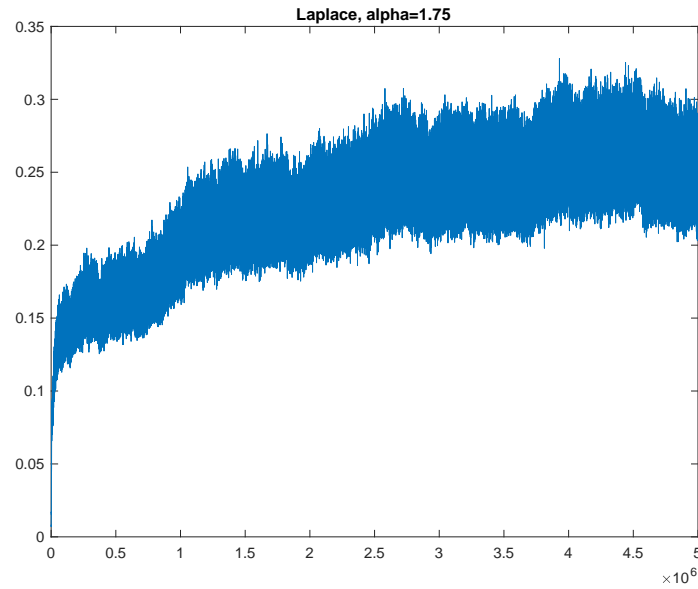
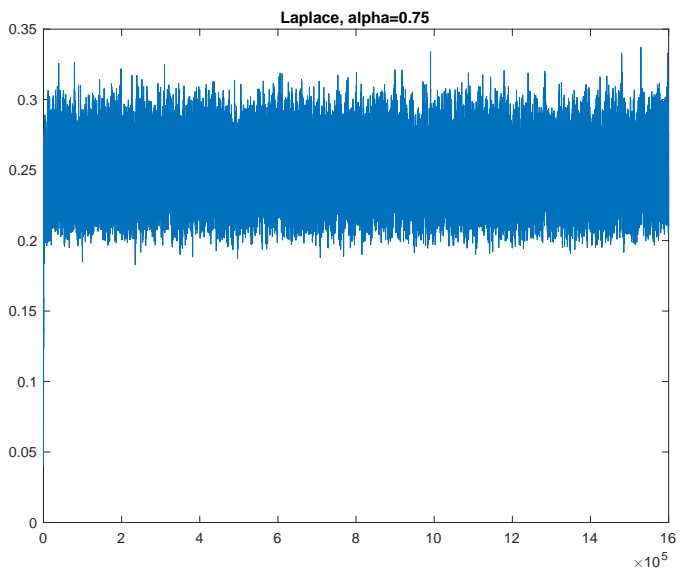
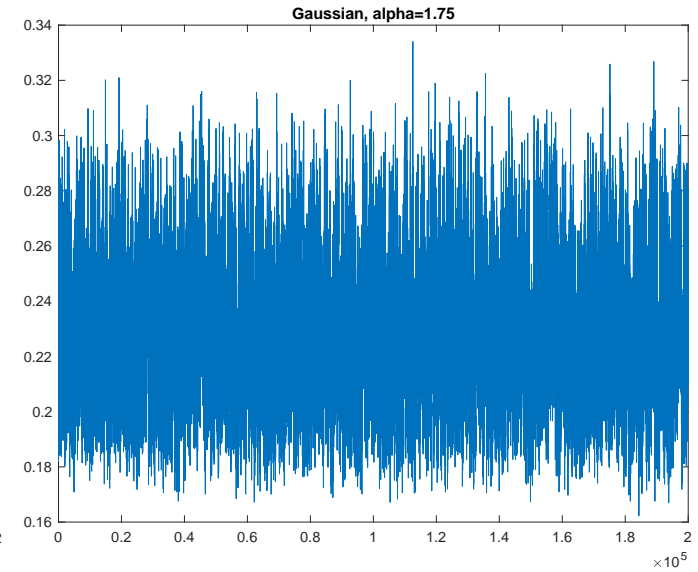
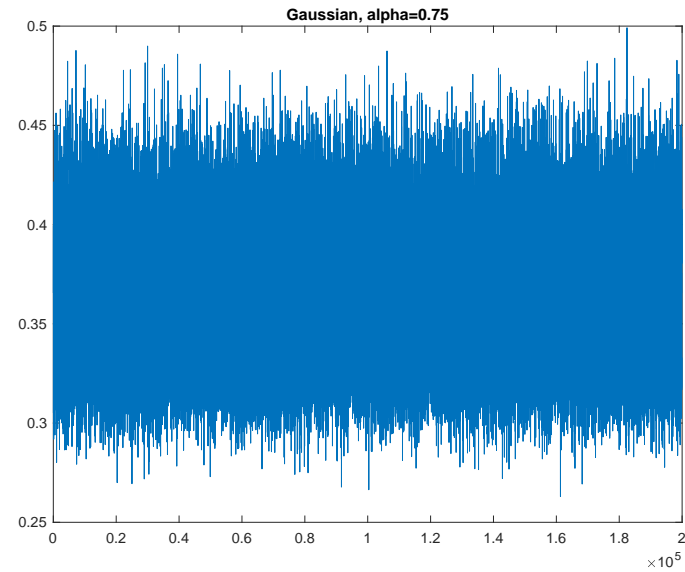
NMR data - Gauss vs Laplace priors



NMR data - Gauss vs Laplace priors - τ -chains



NMR data - Gauss vs Laplace priors - δ -chains



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Summary and open questions




- Laplace priors outperform Gaussian priors over Besov regularity
- Adaptation over Besov spaces (with A. Savva)
- ROC for Bayesian inverse problems with Besov-priors (with S. Wang)
- Sharpness of rates, do we really need scaling and regularity tuning?
- Experimenting with heavier-tailed priors (with I. Castillo)
- For benefit to be realized need better algorithms

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- Laplace priors outperform Gaussian priors over Besov regularity
- Adaptation over Besov spaces (with A. Savva)
- ROC for Bayesian inverse problems with Besov-priors (with S. Wang)
- Sharpness of rates, do we really need scaling and regularity tuning?
- Experimenting with heavier-tailed priors (with I. Castillo)
- For benefit to be realized need better algorithms

THANK YOU!

<http://www.mas.ucy.ac.cy/~sagapi01/>

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